Perpetuities with thin tails, revisited

Paweł Hitczenko* and Jacek Wesołowski†

August 22, 2008

Abstract

We consider the tail behavior of random variables R which are solutions of the distributional equation $R \stackrel{d}{=} Q + MR$, where (Q, M) is independent of R and $|M| \le 1$. Goldie and Grübel showed that the tails of R are no heavier than exponential and that if Q is bounded and M resembles near 1 the uniform distribution, then the tails of R are Poissonian. In this paper we further investigate the connection between the tails of R and the behavior of M near 1. We focus on the special case when Q is constant and M is non–negative, but our results could be easily extended to more general situations.

1 Introduction

In this note we consider a random variable R given by the solution of the stochastic equation

$$R \stackrel{d}{=} Q + MR, \tag{1.1}$$

where (Q, M) are independent of R on the right-hand side. Under suitable assumptions on (Q, M) one can think of R as a limit in distribution of the following iterative scheme

$$R_n = Q_n + M_n R_{n-1}, \quad n \ge 1 \tag{1.2}$$

where R_0 is arbitrary and (Q_n, M_n), $n \ge 1$, are i.i.d. copies of (Q, M), and (Q_n, M_n) is independent of R_{n-1}. Writing out the above recurrence and renumbering the random variables (Q_n, M_n) we see that R may also be defined by

$$R \stackrel{d}{=} \sum_{j=1}^{\infty} Q_j \prod_{k=1}^{j-1} M_k, \tag{1.3}$$

*Department of Mathematics, Drexel University, 3141 Chestnut Street, Philadelphia, PA 19104, phitczenko@math.drexel.edu

†Wydział Matematyki i Nauk Informacyjnych, Politechnika Warszawska, Pl. Politechniki 1, 00-661 Warszawa, Poland, wesolo@mini.pw.edu.pl

provided that the series above converges in distribution. Sufficient conditions for the almost sure convergence are known and have been given by Kesten [13] who also considered a multidimensional case when M is a matrix and Q a vector. For a nice detailed discussion of a one dimensional case we refer to the paper by Vervaat [19]; we only mention briefly here that $\mathbb{E}\log^+|Q| < \infty$ and $\mathbb{E}\log|M| < 0$ suffice for the almost sure convergence of the series in (1.3)

In the form (1.3) R has been studied in insurance mathematics under the name perpetuity. Since schemes like (1.2) are ubiquitous in many areas of applied mathematics, the properties of R have attracted a considerable interest. We refer to [5, 6, 7, 8, 13, 16, 19] and references therein for more information and sample of applications. For examples of more recent work on perpetuities and their applications see [1, 2, 11, 14]. A few additional situations in which perpetuities arise will be mentioned below.

The main focus of research is the tail behavior of R. Kesten [13] showed that if $\mathbb{P}(|M| > 1) > 0$ then R is always heavy–tailed. More precisely, he showed that if there exists a κ such that $\mathbb{E}|M|^\kappa \log^+|M| < \infty$, $\mathbb{E}|Q|^\kappa < \infty$, and $\mathbb{E}|M|^\kappa = 1$ then for some constant C

$$\mathbb{P}(|R| \geq x) \sim Cx^{-\kappa}, \quad \text{as} \quad x \to \infty.$$

Here, and throughout the paper the symbol $f(x) \sim g(x)$ means that the ratio goes to 1 as $x \to \infty$. His result was rediscovered, reproved, and extended by several authors (see [7, 9, 10]). In the complementary case, $\mathbb{P}(|M| \leq 1)$ the picture is much less clear. The main work we are aware of is that of Goldie and Grübel [8] who showed that in that case, the tails are never heavier than exponential and that if M behaves near 1 as a uniform random variable then the tails have Poissonian decay. In their arguments Goldie and Grübel relied on inductive arguments applied to (1.2).

The main purpose of this note is to use systematically their approach to obtain additional information on the links between the behavior of M near 1 and the tail behavior of R. Following Goldie and Grübel (and also customs in large deviation theory) we will be interested in the asymptotics of the logarithm of the tail probability, i.e. $\ln \mathbb{P}(|R| \geq x)$ as $x \to \infty$. Since we are mainly interested in establishing the links between M and R, we will often make additional, but common, assumptions when necessary. For example, we generally assume that Q and M are independent or even that $Q \equiv q$ is non–random. The independence assumption is typically needed only for the lower bounds on the log of the tail probability, the upper bounds are usually obtainable without it. Once the independence of Q and M is assumed the restriction that Q is degenerate does not seem to be a major restriction, but makes some of the arguments more transparent. It is rather the assumption that Q is bounded, which seems to play the more important role. Similarly, we will assume that M and q are non–negative. How the non–negative case differs from the general is relatively well understood (see e.g. arguments in [8, Theorem 2.1, Theorem 3.1, Lemma 5.3]) to see how arguments for non–negative case can be extended to more general situations.

We would like to mention an interesting connection of perpetuities with a subclass of infinitely divisible laws, namely, as was shown by Jurek [12] all self–decomposable random variables (we refer to [12] for the definition) can be represented as perpetuities R given by (1.1) with $0 \leq M \leq 1$. In fact, much more is true, namely, if R is self–decomposable then for *every* random variable $0 \leq M \leq 1$ there exists a random variable Q (typically not bounded) such that (1.1) holds with (Q, M) independent of R on the right–hand side. This curious result seems to be of little help as far as general theory of perpetuities goes. In fact, one can take M to be any constant $M = m \in (0, 1)$ and equally well represent a self–decomposable random variable as a series of weighted iid random variables, with weights forming a geometric progression. Nonetheless, we mention that building on an earlier work of Thorin [18, 17], Bondesson [3] proved a general result which implies, in particular, that all gamma, inverse gamma, Pareto, log–normal, and Weilbull distributions are self–decomposable. Some of these results were obtained earlier by other authors and we refer to Bondesson [3, Section 5] for credits and more examples.

2 General outline

To begin the discussion, assume that $|M| \le 1$. Trivially, if $|Q| \le q$ and $|M|$ is concentrated on a proper subinterval $(0, 1-\delta)$, $\delta > 0$ of $(0,1)$ then the perpetuity R is a random variable whose absolute value is bounded by q/δ and thus has a trivial tail in the sense that $\mathbb{P}(|R| \ge x) = 0$ for $x > q/\delta$. On the other hand if M is not bounded away from 1 then we have the following observation due to Goldie and Grübel:

Proposition 1. *For $\delta \in (0,1)$ let $p_\delta := \mathbb{P}(1-\delta \le M \le 1)$. Then for every such δ and for all $y > 0$ we have*

$$\mathbb{P}(R \ge \frac{q}{\delta}(1-(1-\delta)^y)) \ge p_\delta^y. \tag{2.4}$$

In particular, if for $c \in (0,1)$ and $x > q$ we set

$$\delta = \frac{cq}{x} \quad \textit{and} \quad y = \frac{\ln(1-c)}{\ln(1-cq/x)},$$

then we get that

$$\mathbb{P}(R \ge x) \ge \left(p_{\frac{cq}{x}}\right)^{\frac{\ln(1-c)}{\ln(1-cq/x)}} = \exp\left(\frac{\ln(1-c)}{\ln(1-cq/x)}\ln(p_{\frac{cq}{x}})\right). \tag{2.5}$$

Proof: This was observed by Goldie–Grübel: For a given $\delta > 0$ we let

$$\tau = \tau_\delta = \inf\{n \ge 1 : \; M_n < 1-\delta\}.$$

Then by non–negativity and (1.3), on $\{\tau \ge n\}$ we have

$$R \ge \sum_{k=1}^n q(1-\delta)^{k-1} = \frac{q}{\delta}\left(1-(1-\delta)^n\right).$$

Therefore, for all $n \ge 1$,

$$\mathbb{P}(R \ge \frac{q}{\delta}(1-(1-\delta)^n)) \ge \mathbb{P}(M_k \ge 1-\delta, \; 1 \le k < n) = p_\delta^{n-1}.$$

Hence,

$$\mathbb{P}(R \ge \frac{q}{\delta}(1-(1-\delta)^y)) \ge p_\delta^y, \quad \text{for all } y > 0$$

which proves (2.4); (2.5) follows by a simple calculation. □

It is clear from the above proposition that if p_δ is strictly positive for every $\delta > 0$ then the prepetuity R has non-trivial tails. It is then the behavior of M near 1 that determines the nature of the tails of R. It appears that essentials of such a behavior are shared by a class of equivalent distributions in the following sense.

Let μ and ν be probability distributions on $[0,1]$. For any $\delta \in (0,1)$ we denote $\mu_\delta = \mu((1-\delta,1])$ and $\nu_\delta = \nu((1-\delta,1])$ We say that the distributions μ and ν are equivalent at 1 if

$$\exists\, \varepsilon > 0 \quad \text{and} \quad 0 < d < D < \infty \quad \text{such that}$$
$$\forall \delta \in (0,\varepsilon] : \quad d \le \frac{\mu_\delta}{\nu_\delta} \le D\,. \tag{2.6}$$

As we mentioned earlier, our goal here is to shed some additional light on the relationship between the behavior of the distribution of M in the left neighborhood of 1 and the tails of R. To accomplish that we will develop in a systematic way the approach of Goldie and Grübel. For the upper bound this approach relies on iteration of (1.2) to get a uniform upper bound on the moment generating function of R_n for all $n \ge 1$ and then use exponentiation and Markov inequality to translate this bound into bounds on the tails. We will develop this in the next section, but to give a flavor of this argument we provide the following illustration: Consider (1.1) and assume that Q, M, and R on the right-hand side of (1.1) are independent (that is of course stronger than the usual assumption that (Q,M) are independent of R). Also, assume that $0 \le M \le 1$ and that $m := \mathbb{E}M < 1$. To get an upper bound on the moment generating function $\mathbb{E}e^{zR}$ of R, the principle of what Goldie-Grübel did is the following: for $n \ge 1$ we have

$$\mathbb{E}e^{zR_n} = \mathbb{E}e^{z(Q_n + M_n R_{n-1})} = \mathbb{E}e^{zQ}\mathbb{E}e^{zMR_{n-1}} \le \mathbb{E}e^{zQ}\left\{1 + m\mathbb{E}(e^{zR_{n-1}} - 1)\right\},$$

where in the last step we use the fact that for $s > 0$

$$\mathbb{E}e^{sM} \le \mathbb{E}e^{s\mathrm{Bin}(1,\mathrm{m})} = 1 + m(e^s - 1). \tag{2.7}$$

To set up an induction we seek a function $A(z)$ such that

(i) $Ee^{zR_{n-1}} \le A(z)$, and

(ii) $\mathbb{E}e^{zQ}\left\{1 + m(A(z) - 1)\right\} \le A(z)$.

Solving (ii) gives

$$B(z) := \frac{(1-m)\mathbb{E}e^{zQ}}{1 - m\mathbb{E}e^{zQ}} \le A(z),$$

for z such that $m\mathbb{E}e^{zQ} < 1$. Now, $B(z)$ is recognized as the moment generating function of $\sum_{k=1}^{N} Q_k$ where $N \stackrel{d}{=} \mathrm{Geom}(1-\mathrm{m})$ and is independent of the sequence Q_k, $k \ge 1$. So if we start with any R_0 for which (i) holds with $B(z)$ in place of $A(z)$ then the induction goes through and, under a reasonably weak assumptions on Q, we get an exponential upper bound on the tail of R. In particular if we take $Q \equiv 1$ and $M \stackrel{d}{=} \mathrm{Bin}(1,\mathrm{m})$ then R has moment generating function bounded by that of a geometric random variable and hence sub-exponential tails as was shown already by Goldie and Grübel.

We mention briefly that the sums described by $B(z)$ are yet another example of perpetuities. Sums like these have been studied before and are of interest in renewal theory and risk assessment, for example. For more information and references we refer to [4, 20] where such sums are considered under the name geometric convolutions and geometric random sums, respectively.

As for the lower bound, the best that is available at this point is argument based on Proposition 1. Interestingly, this proposition provides a surprisingly good lower bound. By this we mean the fact that if the upper bound obtained by the above method is constructed carefully so as to be relatively tight, then one can usually obtain a lower bound of a similar strength from Proposition 1. This will be seen in several situations below. It is thus important to understand how to construct a tight upper bound. Although we do not have a general result to that effect, in the last section we will provide heuristic argument that worked well in several situations.

The rest of the paper is organized as follows, in the next section we will discuss an upper bound and in particular, we will state an inequality (see (3.13) below) that is crucial for the inductive argument. In subsequent sections we will illustrate this with several examples. Those include beta(α, β) densities, and what (for the lack of a better name) we call the generalized beta$(1, \beta)$ densities. The reason for considering beta distributions is that one might reasonably hope that they provide a natural parametrization of a behavior of M near 1, which could be translated to the tail behavior of R. This, however, is not the case, since as we will show all beta distributions lead to the same, namely Poissonian, behavior. It turns out that a much more rapid than power–type variability of M at 1 is needed to observe a different tail behavior of R. We will then construct densities for which the logarithm of the tail probability will have power behavior $-x^r$, for $1 < r < \infty$. In the last section we will discuss one more example and we will use it to provide heuristics as to how one might hope to construct an M that would give a particular tail behavior of R.

3 Upper bounds

We begin with the following well–known fact.

Proposition 2. *Suppose that*

$$\mathbb{E}e^{zX} \leq \exp(B\Phi(z)), \tag{3.8}$$

for some function $\Phi : [0, \infty) \to [0, \infty)$*,* $B > 0$ *and all* $z > 0$*. Then*

$$\mathbb{P}(X \geq x) \leq e^{-\Phi^*(x))}, \tag{3.9}$$

where $\Phi^* = \Phi^*_B$ *is defined by*

$$\Phi^*(x) = \sup\{zx - B\Phi(z) : z > 0\}. \tag{3.10}$$

Note that if Φ is an Orlicz function (a convex, continuous, non–decreasing function, such that $\Phi(0) = 0$ and $\Phi(t) \to \infty$ as $t \to \infty$) then Φ^* is just a complementary function to Φ.

Proof: This is well–known; by the usual exponentiation and Markov's inequality we have

$$\mathbb{P}(X \geq x) = \mathbb{P}(e^{zX} \geq e^{zx}) \leq e^{-zx}\mathbb{E}e^{zX} \leq e^{-zx}e^{B\Phi(z)} = e^{-(zx - B\Phi(z))}.$$

Since the right–hand side may be minimized over z we obtain (3.9) as required. $\square$

One can obtain a bound on the moment generating function of R using the fact that it is a limit in distribution of the iterative procedure (1.2) and verifying (3.8) for every R_n. In the case $Q_n \equiv q$ (1.2) takes the form

$$R_n \stackrel{d}{=} q + M_n R_{n-1}, \tag{3.11}$$

where M_n is a copy of M independent of R_{n-1}. To argue inductively, suppose that for some $B > 0$

$$\mathbb{E}e^{zR_{n-1}} \leq \exp(B\Phi(z)), \quad z > 0. \tag{3.12}$$

Then by (3.11) and (3.12) applied conditionally on M_n we have

$$\mathbb{E}e^{zR_n} = e^{qz}\mathbb{E}e^{zM_nR_{n-1}} \leq e^{qz}\mathbb{E}e^{B\Phi(zM_n)}.$$

The inductive step will be complete once we show that

$$e^{qz}\mathbb{E}e^{B\Phi(zM)} \leq e^{B\Phi(z)}.$$

In terms of the distribution μ of M, the above inequality reads

$$e^{qz}\int_0^1 e^{B\Phi(zt)}\mu(dt) \leq e^{B\Phi(z)}. \tag{3.13}$$

Once this inequality is established, the induction is complete as one can start with arbitrary random variable R_0, so in particular we can ensure that (3.12) holds for R_0. The above inequality is crucial for establishing the upper bound.

We will be interested in the tail bounds for large values of x. We assume that Φ is non-degenerate ($\Phi(t) \neq 0$ for $t \neq 0$) and satisfies $\Phi(t)/t \to \infty$ as $t \to \infty$ (i.e. Φ is an N–function in the language of [15]). Then Φ^* has the same properties and it follows directly from the definition (3.10) that as $x \to \infty$ the supremum in (3.10) is attained at $z \to \infty$. This means that it suffices that (3.8) and thus (3.13) hold only for large values of z. Thus, we have the following consequence of the above discussion:

Proposition 3. *Let R be given by (1.1) with $Q \equiv q$. Suppose that M has a density such that there exist $B > 0$ and z_0 such that (3.13) is satisfied for all $z \geq z_0$. Then*

$$\limsup_{x\to\infty} \frac{\ln \mathbb{P}(R \geq x)}{\Phi_B^*(x)} \leq -1. \tag{3.14}$$

4 Beta distributions

As earlier we will denote by μ the distribution of M. Goldie–Grübel [8, Theorem 3.1] showed that if Q is bounded and μ and the uniform distribution on $[0, 1]$ are equivalent at 1 in the sense of (2.6) then the resulting perpetuity has Poissonian tails, that is

$$\lim_{x\to\infty} \frac{\ln \mathbb{P}(R \geq x)}{x \ln x} = -\frac{1}{q}.$$

Note that uniform and beta $\beta(\alpha, 1)$ distributions are equivalent at 1. One might reasonably hope that considering other values of the second parameter of the beta distribution might lead to a different tail behavior of R but this is not the case. As we show below *any* M whose distribution is equivalent at 1 to a measure with polynomial density at 1 leads to the Poissonian tails of R.

Theorem 4. *Let the distribution of M and the beta(α,β) distribution be equivalent at 1. Assume that $Q \equiv q > 0$. Then*

$$\lim_{x\to\infty} \frac{\ln \mathbb{P}(R \geq x)}{x \ln x} = -\frac{\beta}{q}.$$

Proof: Note that all beta distributions with the same β parameter and different α parameters are equivalent in the sense of (2.6). Consequently, we assume for convenience that $\alpha = 1$ so that we consider the beta distribution with the density

$$f(t) = \beta(1-t)^{\beta-1}, \quad 0 < t < 1 ,$$

which is equivalent to the distribution of M at 1.

We show that regardless of the value of $\beta > 0$ the tails of the resulting perpetuities are Poissonian. To get an upper bound we verify that (3.13) holds with $\Phi(z) = e^{bz}$ for a suitable constant b and some $B > 0$. Once this is done, it follows from the discussion in the previous section that

$$\ln \mathbb{P}(R \geq x) \leq -\frac{x}{b} \ln\left(\frac{x}{Bbe}\right) = -\frac{1}{b} x \left(\ln x - \ln(Bbe)\right).$$

which implies that

$$\limsup_{x\to\infty} \frac{\ln \mathbb{P}(R \geq x)}{x \ln x} \leq -\frac{1}{b}. \tag{4.15}$$

Thus we are to show that for sufficiently large $z > 0$

$$e^{qz} \int_0^1 \exp(Be^{bzt})\mu(dt) \leq \exp(Be^{bz}), \tag{4.16}$$

for some positive constant B and $b = q/\beta$. To that end take an ε for which (2.6) holds with ν being a beta$(1, \beta)$ distribution. Assume a t_0 is chosen so that $t_0 > 1 - \varepsilon$. We split the integral on the left–hand side as

$$e^{qz} \int_0^{t_0} \exp(Be^{bzt})\mu(dt) + e^{qz} \int_{t_0}^1 \exp(Be^{bzt})\mu(dt).$$

The second term, through (2.6) is bounded by

$$De^{qz} \exp(Be^{bz})\beta \int_{t_0}^1 (1-t)^{\beta-1} dt = De^{qz} \exp(Be^{bz})(1-t_0)^\beta.$$

Pick $t_0 = t_0(z) > 1 - \varepsilon$ so that

$$\rho := De^{qz}(1-t_0)^\beta < 1. \tag{4.17}$$

In order to establish (4.16) we are to show that

$$e^{qz} \int_0^{t_0} \exp(Be^{bzt})\mu(dt) \leq (1-\rho) \exp(Be^{bz}).$$

It follows from (4.17) that

$$t_0 = 1 - e^{-qz/\beta}(\rho/D)^{1/\beta},$$

and thus for sufficiently large z we have that $t_0 > 1 - \varepsilon$. Hence, the left–hand side above, by (2.6) again, is bounded by

$$e^{qz}\exp(Be^{bzt_0})\mu(0,t_0) \le e^{qz}\exp(Be^{bzt_0})\left(1 - \frac{d}{D}\rho e^{-qz}\right),$$

and we want this to be less or equal than $(1-\rho)\exp(Be^{bz})$. Divide both sides by $\exp(Be^{qz})$ so that the inequality to be proved reads

$$e^{qz}\exp\left(Be^{bzt_0} - Be^{bz}\right)\left(1 - \frac{d}{D}\rho e^{-qz}\right) \le 1 - \rho.$$

We drop the factor $1 - \frac{d}{D}\rho e^{-qz}$ on the left and look at the exponent. It is

$$qz + Be^{bz(1-e^{-qz/\beta}(\rho/D)^{1/\beta})} - Be^{bz} = qz + Be^{bz}\left(e^{-bze^{-qz/\beta}(\rho/D)^{1/\beta}} - 1\right).$$

Set $b := q/\beta$. Since $\rho/D < 1$ we have $bze^{-qz/\beta}(\rho/D)^{1/\beta} = bze^{-bz}(\rho/D)^{1/\beta} < bze^{-bz} \le e^{-1} < \ln 2$. Since $e^{-u} - 1 \le -u/2$ for $0 < u < \ln 2$ we see that the expression above is bounded by

$$qz - Bbz\rho^{1/\beta}e^{bz}e^{-bz}/2 = qz\left(1 - \frac{B\rho^{1/\beta}}{2\beta}\right),$$

and it is clear that

$$e^{qz}\exp\left(Be^{bzt_0} - Be^{bz}\right) \le \exp\left(qz(1 - \frac{B\rho^{1/\beta}}{2\beta})\right),$$

can be made arbitrarily small by increasing B if necessary. In particular, we can ensure that it is less than $1-\rho$ for all z not too close to 0. Thus, (4.15) is proved with $b = q/\beta$.

To get the matching lower bound note that using again instead of M the equivalent law $\mathrm{beta}(1,\beta)$ with the cdf $F(t) = 1 - (1-t)^\beta$ we have

$$\nu_\delta = 1 - F(1-\delta) = \delta^\beta.$$

Thus, by (2.6)

$$\begin{aligned}\mathbb{P}(R \ge x) &\ge \left(d\frac{cq}{x}\right)^{\beta\frac{\ln(1-c)}{\ln(1-cq/x)}} = \exp\left(-\beta\frac{\ln(1-c)}{\ln(1-cq/x)}(\ln x - \ln(dcq))\right)\\ &= \exp\left(\beta\frac{\ln(1-c)}{cq}(x\ln x)(1+o(1))\right).\end{aligned}$$

Hence, by letting $c \to 0_+$ we get that

$$\liminf_{x\to\infty}\frac{\ln\mathbb{P}(R \ge x)}{x\ln x} \ge -\frac{\beta}{q}.$$

□

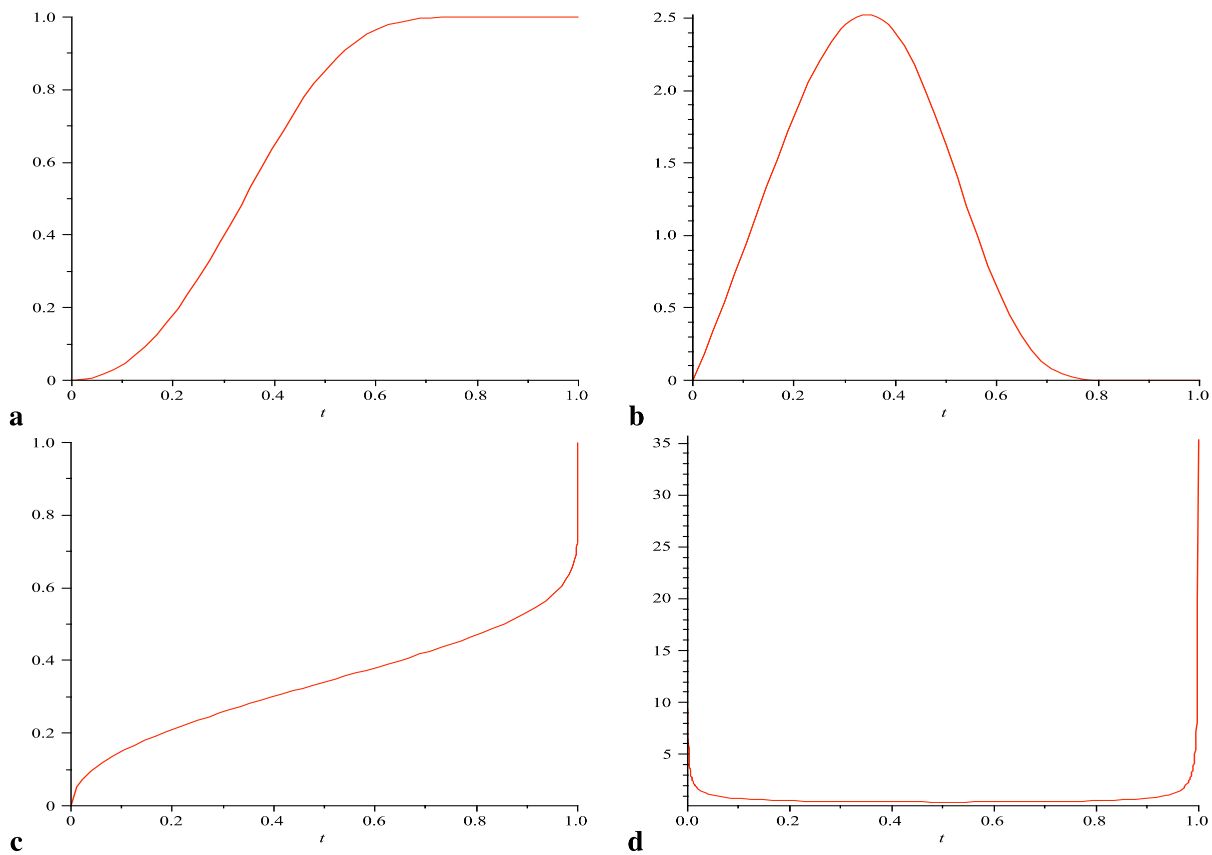

Figure 1: (a) The distribution $F_{4,2}$, (b) its density, (c) its inverse $F_{0.5,0.5}$, and (d) its density.

5 Generalized beta$(1, \beta)$ distributions

In this section we consider M's whose distributions are equivalent in the sense (2.6) to distribution function given by

$$F(s) = F_{\beta,\eta}(s) = 1 - e^{-\beta(-\ln(1-s))^{\eta}}, \quad 0 < s < 1, \quad \beta, \eta > 0. \tag{5.18}$$

It is elementary to verify that $F_{\beta,\eta}$ is indeed a distribution function which is strictly increasing on $(0, 1)$. Furthermore, $F_{\beta,1}$ is the distribution of a beta$(1, \beta)$ random variable discussed in the previous section. The family $F_{\beta,\eta}$ has the following property

$$F_{\beta,\eta}^{-1} = F_{\beta^{-1/\eta},\eta^{-1}},$$

as can be easily verified by a direct calculation. Pictures of a few such distributions with various parameters are given in Figures 1–2.

For R generated with M's with distributions equivalent to the above distribution function the following extension of Theorem 4 holds

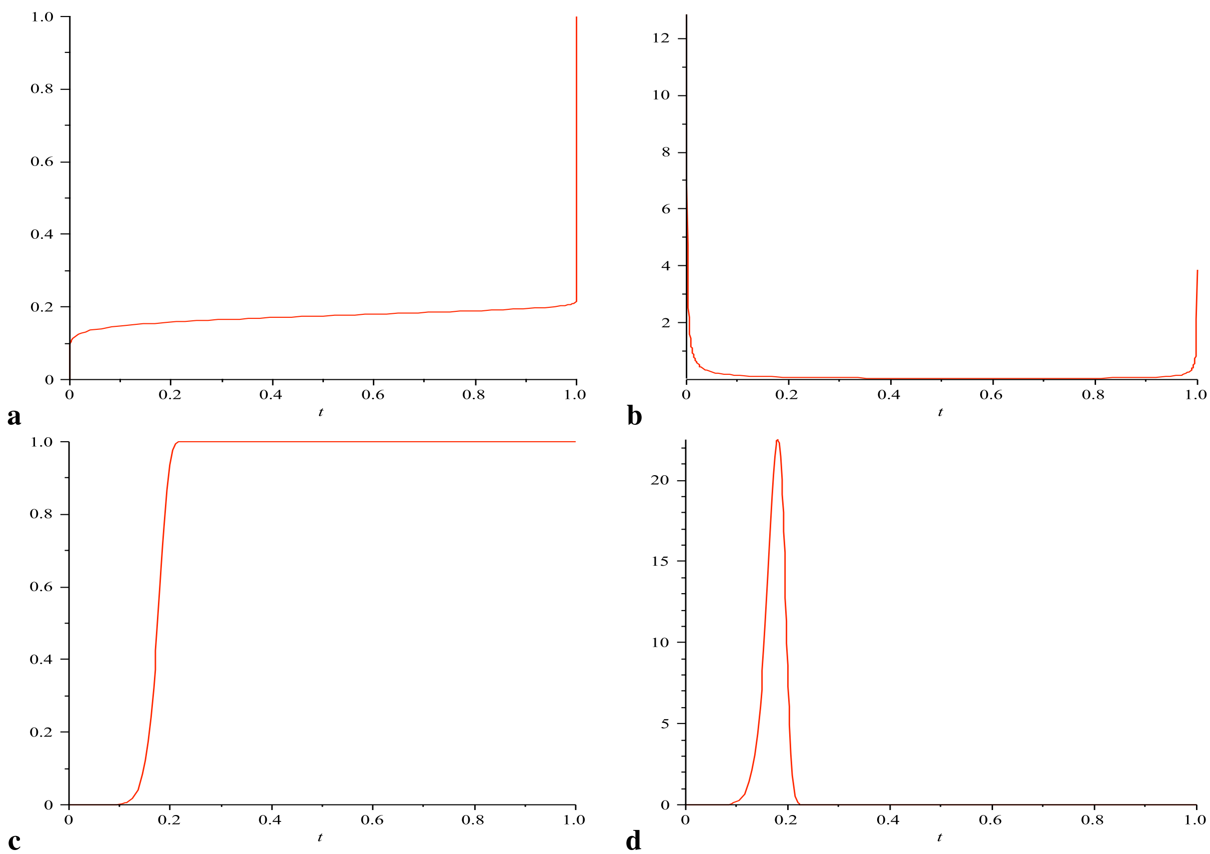

Figure 2: (a) The distribution $F_{0.2,0.1}$, (b) its density, (c) its inverse $F_{5^{10},10}$, and (d) its density.

Theorem 5. *Let* (R_n) *given by (3.11) where* $q > 0$ *and* M *has the distribution equivalent to the distribution function (5.18) for some* $\beta, \eta > 0$*. Let* R *be a limit in distribution of* (R_n)*. Then*

$$\lim_{x\to\infty} \frac{\ln \mathbb{P}(R \geq x)}{x(\ln x)^\eta} = -\frac{\beta}{q}.$$

Proof: For the upper bound we will show that R satisfies Proposition 3 with $\Phi(z) = \exp(bz^{1/\eta})$ for b's in a certain range. For this Φ we have

$$\Phi^*(x) \geq x\left(\left(\frac{\ln x}{b}\right)^\eta - B\right)$$

which can be seen by using $\Phi_B^*(x) \geq xz_0 - Be^{bz_0^{1/\eta}}$ with $z_0 = b^{-\eta}(\ln x)^\eta$. It follows that

$$\limsup_{x\to\infty} \frac{\ln \mathbb{P}(R \geq x)}{x(\ln x)^\eta} \leq -\frac{1}{b^\eta}. \tag{5.19}$$

To verify (3.13) we will use the same argument as before; with $\Phi(z) = \exp(bz^{1/\eta})$ it becomes

$$e^{qz}\int_0^1 \exp(Be^{b(zt)^{1/\eta}})\mu(dt) \leq \exp(Be^{bz^{1/\eta}}),$$

where μ is the distribution of the rv M and b and B are positive constants. Splitting the left–hand side, with $t_0 > 1-\varepsilon$ as before we have

$$\beta e^{qz}\int_0^{t_0} \exp(Be^{b(zt)^{1/\eta}})\mu(dt) + e^{qz}\int_{t_0}^1 \exp(Be^{b(zt)^{1/\eta}})\mu(dt).$$

By (2.6) the second term is bounded by

$$De^{qz}\exp\left(Be^{bz^{1/\eta}}\right)(1 - F_{\beta,\eta}(t_0)).$$

Choose t_0 so that $\rho := De^{qz}(1-F(t_0)) < 1$. Then

$$\begin{aligned} t_0 &= F_{\beta,\eta}^{-1}(1-\rho e^{-qz}/D) = F_{\beta^{-1/\eta},\eta^{-1}}(1-\rho e^{-qz}/D) = 1 - \exp\left(-\beta^{-1/\eta}(-\ln(\rho e^{-qz}/D))^{1/\eta}\right) \\ &= 1 - \exp\left(-\left(\frac{qz}{\beta}\right)^{1/\eta}\left(1 - \frac{\ln(\rho/D)}{qz}\right)^{1/\eta}\right), \end{aligned}$$

and for z sufficiently large it follows that $t_0 > 1-\varepsilon$. Now we are to prove that

$$e^{qz}\exp\left(Be^{bz^{1/\eta}t_0^{1/\eta}}\right)\mu(0,t_0) \leq (1-\rho)\exp\left(Be^{bz^{1/\eta}}\right).$$

By the first part of (2.6), it is enough to show that

$$e^{qz}\exp\left(Be^{bz^{1/\eta}}\left(e^{-bz^{1/\eta}(1-t_0^{1/\eta})} - 1\right)\right)\left(1 - \frac{d\rho}{D}e^{-qz-Be^{bz^{1/\eta}}}\right) \leq 1-\rho. \tag{5.20}$$

We drop the last factor on the left–hand side as it is less that 1. For t_0 as above $z^{1/\eta}(1-t_0^{1/\eta})$ is close to 0 for z sufficiently large, so that using approximations $e^{-x}-1\sim -x$ and then $1-(1-x)^\eta\sim x/\eta$, both valid for x close to 0 we see that the exponent on the left–hand side for z sufficiently large, is

$$\begin{aligned} qz + Be^{bz^{1/\eta}}\left(e^{-bz^{1/\eta}(1-t_0^{1/\eta})}-1\right) &\sim qz - Bbz^{1/\eta}e^{bz^{1/\eta}}(1-t_0^{1/\eta}) \\ &\sim qz-\frac{Bb}{\eta}z^{1/\eta}\exp\left(z^{1/\eta}\left\{b-\left(\frac{q}{\beta}\right)^{1/\eta}\left(1-\frac{\ln\rho/D}{qz}\right)^{1/\eta}\right\}\right) \\ &\sim qz-\frac{Bb}{\eta}z^{1/\eta}\exp\left(z^{1/\eta}\left\{b-(q/\beta)^{1/\eta}\right\}\right). \end{aligned}$$

For $b>(q/\beta)^{1/\eta}$ the second term grows faster than linearly in z, so that as long as z is not too close to 0 it can be made arbitrarily larger than qz. Thus, (5.20) follows. Furthermore, letting $b\to(q/\beta)_+^{1/\eta}$ in (5.19) we obtain that

$$\limsup_{x\to\infty}\frac{\ln\mathbb{P}(R\ge x)}{x(\ln x)^\eta}\le-\frac{\beta}{q}. \tag{5.21}$$

To get a lower bound note that, using instead of the distribution of M the equivalent cdf $F_{\beta,\eta}$, on noting that

$$1-F_{\beta,\eta}(1-cq/x)=\exp\left(-\beta\left(-\ln(cq/x)\right)^\eta\right)=\exp\left(-\beta(\ln x-\ln(cq))^\eta\right)$$

we get for large x

$$\begin{aligned} \mathbb{P}(R\ge x)\ge\left(d(1-F_{\beta,\eta}(1-cq/x))\right)^{\frac{\ln(1-c)}{\ln(1-cq/x)}} &= \exp\left(-\frac{\ln(1-c)}{\ln(1-cq/x)}\beta[(\ln x-\ln(cq))^\eta+\ln(d)]\right) \\ &= \exp\left(\frac{\beta\ln(1-c)}{cq}x(\ln x)^\eta(1-o(1))\right). \end{aligned}$$

Upon letting $c\to 0_+$ it implies that

$$\liminf_{x\to\infty}\frac{\ln\mathbb{P}(R\ge x)}{x(\ln x)^\eta}\ge-\frac{\beta}{q}\ .$$

Combining this with (5.21) completes the proof. □

6 Weilbull–like tails

In this section we explicitly construct M's that will lead to a rather different tail behavior of R than discussed in the previous sections. As we will see a much more rapid variability of M near 1 is needed to obtain a lighter tail behavior of R. More specifically, we prove the following theorem.

Theorem 6. *Let $1<r<\infty$. Let the distribution of M be (2.6) equivalent to the distribution ν with the density*

$$f_\nu(t)\ \propto\ t^{r-1}e^{-\frac{1}{(1-t^r)^{1/(r-1)}}}I_{(0,1)}(t). \tag{6.22}$$

Then for the perpetuity R given by (1.3) with $Q \equiv q$ there are constants c_1, c_2 such that

$$-\infty < c_1 \le \liminf_{x\to\infty} \frac{\ln \mathbb{P}(R \ge x)}{(x/q)^r} \le \limsup_{x\to\infty} \frac{\ln \mathbb{P}(R \ge x)}{(x/q)^r} \le c_2 < 0.$$

Proof: For $1 < r < \infty$ let r^* be given by

$$\frac{1}{r} + \frac{1}{r^*} = 1.$$

The role of r and r^* are symmetric and for notational convenience we will prove the above inequalities for r^* rather than r. Suppose we prove that for M the condition (3.12) holds for all $n \ge 1$ with $\Phi(z) = z^r$ and some $B > 0$. Then by elementary calculation $\Phi^*(x) = \frac{x^{r^*}}{r^*(Br)^{1/(r-1)}}$, so that,

$$\mathbb{P}(R \ge x) \le \exp\left(-\frac{x^{r^*}}{r^*(Br)^{1/(r-1)}}\right), \tag{6.23}$$

and this would give the claimed behavior of the logarithm of the tail probability of R.

To establish (3.12) via inductive argument, we need to verify that (3.13) holds in the present situation, that is, we want to show that for z sufficiently large

$$e^{qz} \int_0^1 e^{B(zt)^r} \mu(dt) \le e^{Bz^r} .$$

Take $\varepsilon > 0$ given by (2.6) where ν has density given by (6.22) and consider $\delta \in (0, \varepsilon)$. Then the left hand side of the above inequality is less than

$$e^{qz} e^{Bz^r(1-\delta)^r} + e^{qz} \int_{1-\delta}^1 e^{B(zt)^r} \mu(dt) \le e^{qz} e^{Bz^r(1-\delta)^r} + De^{qz} \int_{1-\delta}^1 e^{B(zt)^r} \nu(dt) .$$

Consequently, we have to show that

$$e^{qz - Bz^r(1-(1-\delta)^r)} + De^{qz-Bz^r} \int_{1-\delta}^1 e^{B(zt)^r} f_\nu(t)dt \le 1 . \tag{6.24}$$

Note that because $r > 1$ and $0 < \delta < 1$, the first term can be made arbitrarily small for $z \ge z_0$ sufficiently large. We thus concentrate on the second term. The following argument will not only complete justification of (6.24) but will also indicate how one would be led to a reasonable choice of f_ν if it were unknown. We would want to construct a density f_ν on $(0, 1)$ for which (6.24) holds. To this end suppose for now that the density f_ν were of the form

$$f_\nu(t) = rt^{r-1} g(t^r).$$

Upon changing variables to $s = t^r$ the second term in (6.24) becomes

$$De^{qz-Bz^r} \int_{(1-\delta)^r}^1 e^{Bz^r s} g(s)ds = De^{qz} \int_{(1-\delta)^r}^1 e^{-Bz^r(1-s)} g(s)ds.$$

Setting $w = 1 - s$ gives

$$De^{qz} \int_0^{1-(1-\delta)^r} e^{-Bz^r w} g(1-w)dw. \tag{6.25}$$

We now let

$$g(1-w) := Ke^{-1/w^\gamma},$$

where γ is to be chosen momentarily and $K = K(\gamma)$ is set so that

$$K^{-1} = \int_0^1 e^{-1/w^\gamma} dw.$$

Then (6.25) becomes

$$KDe^{qz} \int_0^{1-(1-\delta)^r} e^{-Bz^r w} e^{-1/w^\gamma} dw. \tag{6.26}$$

The integrand is

$$\exp\left(-(Bz^r w + \frac{1}{w^\gamma})\right).$$

Since the function

$$w \to Bz^r w + \frac{1}{w^\gamma},$$

has a minimum at $(\gamma/(Bz^r))^{1/(\gamma+1)}$ whose value is

$$(Bz^r)^{\frac{\gamma}{\gamma+1}} (\gamma^{\frac{1}{\gamma+1}} + \gamma^{-\frac{\gamma}{\gamma+1}}) = B^{\frac{\gamma}{\gamma+1}} z^{r\frac{\gamma}{\gamma+1}} \frac{\gamma+1}{\gamma^{\gamma/(\gamma+1)}},$$

the quantity (6.26) is no more than

$$KD \exp\left(zq - z^{r\frac{\gamma}{\gamma+1}} B^{\frac{\gamma}{\gamma+1}} \frac{\gamma+1}{\gamma^{\gamma/(\gamma+1)}}\right),$$

which upon setting

$$r\frac{\gamma}{\gamma+1} = 1 \quad \text{i.e.} \quad \gamma = \frac{1}{r-1},$$

becomes

$$KD \exp\left\{z\left(q - B^{1/r}\frac{r}{(r-1)^{(r-1)/r}}\right)\right\}.$$

It is now clear that if

$$B = A^r \left(\frac{q}{r}\right)^r (r-1)^{r-1}, \tag{6.27}$$

where $A > 1$ might depend on r, then $q - B^{1/r} r/(r-1)^{\frac{r-1}{r}} = q(1-A) < 0$. Therefore, for $z \geq z_0$ we obtain further

$$KD \exp\left\{z\left(q - B^{1/r}\frac{r}{(r-1)^{(r-1)/r}}\right)\right\} \leq KDe^{-z_0 q(A-1)}.$$

Thus we conclude that for $z \geq z_0$ the left–hand side of (6.24) is bounded by

$$e^{-z_0(B(1-(1-\delta)^r)z_0^{r-1}-q)} + KDe^{-z_0 q(A-1)}.$$

Since the value of this expression can be made smaller than 1 by choosing z_0 sufficiently large, (6.24) follows.

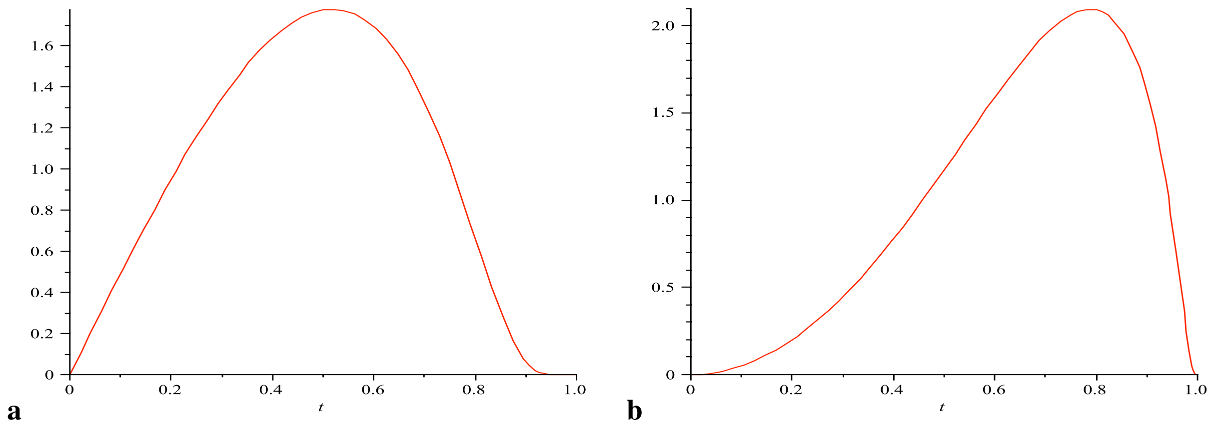

Figure 3: The density (6.22) for (a) $r = 2$ and (b) $r = 3$.

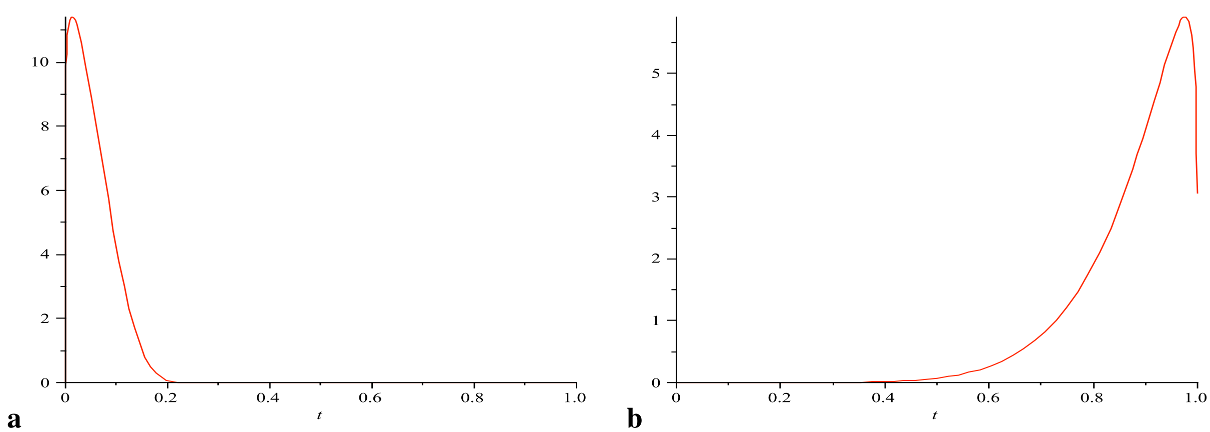

Figure 4: The density (6.22) for (a) $r = 1.1$ and (b) $r = 8$.

Reversing the steps, we obtain the expression for the density f_ν given in (6.22) with the normalizing constant K_r given by

$$K_r^{-1} = \frac{1}{r}\int_0^1 \exp\left(-\frac{1}{v^{1/(r-1)}}\right) dv.$$

Finally, putting the value of B given in (6.27) into (6.23) we obtain

$$\mathbb{P}(R \geq x) \leq \exp\left(-\left(\frac{x}{q}\right)^{r^*} \frac{1}{A^{r/(r-1)}}\right),$$

which implies that

$$\limsup_{x\to\infty} \frac{\ln \mathbb{P}(R \geq x)}{(x/q)^{r^*}} \leq -A^{-r/(r-1)}.$$

To get a lower bound for $\mathbb{P}(R \geq x)$ we choose $\delta \in (0, \varepsilon)$ as in (2.6). Then, upon passing to the equivalent measure with density (6.22) we have

$$p_\delta \geq dK_r \int_{1-\delta}^1 t^{r-1} \exp\left(-\frac{1}{(1-t^r)^{1/(r-1)}}\right) dt.$$

Changing variables to $v = (1-t^r)^{-1/(r-1)}$ yields

$$p_\delta \geq K \int_{(1-(1-\delta)^r)^{-1/(r-1)}}^\infty \frac{e^{-v}}{v^r} dv,$$

for some constant K whose value is irrelevant. Since for large v_0, $\int_{v_0}^\infty \frac{e^{-v}}{v^r} dv$ is comparable to e^{-v_0}/v_0^r we get, up to an unimportant constant

$$(1-(1-\delta)^r)^{r/(r-1)} \exp\left(-\frac{1}{(1-(1-\delta)^r)^{1/(r-1)}}\right)$$

as the lower bound for p_δ. Hence, up to unimportant additive terms

$$\ln p_\delta \geq \frac{r}{r-1} \ln\left(1-(1-\delta)^r\right) - \frac{1}{(1-(1-\delta)^r)^{1/(r-1)}} \sim -\frac{1}{(1-(1-\delta)^r)^{1/(r-1)}},$$

as the second term above is of dominant order for $\delta \to 0$. For small δ we have

$$1-(1-\delta)^r = 1-\exp\left(r\ln(1-\delta)\right) \sim -r\ln(1-\delta),$$

so that upon replacing δ by cq/x we get that, asymptotically

$$\ln p_{cq/x} \geq -\frac{1}{(-r\ln(1-cq/x))^{1/(r-1)}} \sim -\left(\frac{x}{cqr}\right)^{\frac{1}{r-1}}.$$

Combining this with (2.5) we get that, asymptotically,

$$\begin{aligned} \ln \mathbb{P}(R \geq x) &\geq \frac{\ln(1-c)}{\ln(1-cq/x)} \left(- \left(\frac{x}{cqr} \right)^{1/(r-1)} \right) \sim \frac{x \ln(1-c)}{cq} \left(\frac{x}{cqr} \right)^{1/(r-1)} \\ &= \left(\frac{x}{q} \right)^{r^*} \cdot \frac{\ln(1-c)}{(cr^{1/r})^{r^*}}. \end{aligned}$$

It follows that

$$\liminf_{x \to \infty} \frac{\ln \mathbb{P}(R \geq x)}{(x/q)^{r^*}} \geq \frac{C}{r^{1/(r-1)}}, \quad \text{where} \quad C = \frac{\ln(1-c)}{c^{r^*}} < 0.$$

□

Remarks:

(i) The maximal value of $C/r^{1/(r-1)}$ is obtained by setting $c = c_0$ where c_0 is the unique solution of the equation

$$\frac{1}{1-c} + r^* \cdot \frac{\ln(1-c)}{c} = 0.$$

The uniqueness of the solution is elementary as the function

$$h(c) := \frac{\ln(1-c)}{c^{r^*}}$$

approaches $-\infty$ as $c \to 0_+$ or $c \to 1_-$ and

$$h'(c) = -c^{-r^*} \left(\frac{1}{1-c} + r^* \cdot \frac{\ln(1-c)}{c} \right).$$

The expression in the parentheses, upon letting $y = 1/(1-c)$, $y > 1$, becomes

$$y - r^* \cdot \frac{\ln y}{(y-1)/y} = y \left(1 - r^* \cdot \frac{\ln y}{y-1} \right).$$

Since $\frac{\ln y}{y-1}$ is decreasing for $y > 1$, approaches 1 as $y \to 1_+$ and 0 as $y \to \infty$ we see that $h'(c)$ has exactly one sign change (from positive to negative) on $(0,1)$ and that this change occurs at c_0 such that

$$\frac{1}{1-c_0} + r^* \cdot \frac{\ln(1-c_0)}{c_0} = 0.$$

While the above equation does not have in general the closed form solution for c_0 as a function of r (or r^*), the asymptotic behavior of the constant $C/r^{1/(r-1)}$ as r goes to 0 or ∞ can be traced down. Since

$$r^* = -\frac{c_0}{(1-c_0)\ln(1-c_0)},$$

as $r \to \infty$ (and thus $r^* \to 1_+$) we must have $c_0 \to 0_+$ at the rate $1 - c_0 \sim 1/r^*$. But then $c_0 \sim 1 - 1/r^* = 1/r$ and thus

$$\frac{\ln(1-c_0)}{r^{1/(r-1)} c_0^{r^*}} \sim \frac{\ln(1-1/r)}{r^{1/(r-1)}(1/r)^{r/(r-1)}} = r \ln(1-1/r) \to -1, \quad \text{as} \quad r \to \infty.$$

Similarly, if $r \to 1_+$ then $c_0 \to 1_-$ in such a way that $1 - c_0 \sim 1/(r^* \ln r^*)$. Then

$$\frac{\ln(1-c_0)}{r^{1/(r-1)} c_0^{r^*}} \sim \frac{-\ln(r^* \ln r^*)}{r^{1/(r-1)}(1-1/(r^* \ln r^*))^{r^*}} \sim \frac{-\ln(r^* \ln r^*)}{e}$$

since

$$r^{1/(r-1)} = \left(1 + \frac{1}{1/(r-1)}\right)^{1/(r-1)} \to e, \quad \text{and} \quad \left(1 - \frac{1}{r^* \ln r^*}\right)^{r^*} \to 1 \quad \text{as} \quad r \to 1_+.$$

(ii) It might appear from the argument that the form of density (6.22) was just guessed. While it is true that originally this was the case, there is a heuristic argument which would suggest the same choice. We will explain this heuristics in the next section on a different example, but we would like to mention that following it in the present situation would essentially lead to density given by (6.22).

7 Further example

In this section we present one more example of perpetuity that will have extremely thin tails. Specifically, we will show

Proposition 7. *There exist densities f_M for which the perpetuity defined by (3.11) satisifes:*

$$\forall\ B > q \quad \limsup_{x\to\infty} \frac{\ln \mathbb{P}(R \geq x)}{B \exp(x/B)} \leq -\frac{1}{e}, \tag{7.28}$$

and

$$\forall\ B < q \quad \liminf_{x\to\infty} \frac{\ln \mathbb{P}(R \geq x)}{B \exp(x/B)} \geq \frac{\ln(1-B/q)}{B}. \tag{7.29}$$

Proof: We consider the case $\Phi(z) = z \ln z$ and we will show that Proposition 3 holds for all $B > q$. It will then follow that for all such B

$$\mathbb{P}(R \geq x) \leq \exp(-B \exp(\frac{x}{B} - 1)), \tag{7.30}$$

which will imply (7.28). We will then construct a density of M which, on one hand will guarantee (7.30) and, on the other hand, ensure that p_δ is sufficiently large so that the argument based on Proposition 1 will give (7.29).

To carry out the details of that plan we are to construct a density f_M for which

$$e^{qz} \int_0^1 e^{Bzt \ln(zt)} f_M(t) dt \leq e^{Bz \ln z}.$$

This is equivalent to

$$e^{qz} \int_0^1 e^{-B(1-t)z \ln z} t^{Btz} f_M(t) dt \leq 1,$$

and it is enough to construct an f_M for which

$$e^{qz}\int_0^1 e^{-B(1-t)z\ln z} f_M(t)dt = e^{qz}\int_0^1 e^{-Btz\ln z} f_M(1-t)dt \le 1.$$

We now set $f_M(1-t) = K\exp(-h(t))$, where h is a non–negative function and $K^{-1} = \int_0^1 \exp(-h(t))dt$. The inequality to be established becomes

$$e^{qz}\int_0^1 e^{-Btz\ln z - h(t)}dt \le \int_0^1 e^{-h(t)}dt. \tag{7.31}$$

One is guided to a reasonable choice of h by the following heuristics. Suppose h is differentiable and chosen so that

$$Btz\ln z + h(t) \tag{7.32}$$

is minimized at its critical point $t = t_z \in (0,1)$ which thus satisfies

$$Bz\ln z + h'(t_z) = 0. \tag{7.33}$$

Then the left–hand side of (7.31) is no more than

$$\exp\left(qz - Bt_z z\ln z - h(t_z)\right) \le \exp\left(z(q - Bt_z\ln z)\right).$$

Since we must be able to make it arbitrarily negative (by increasing B if necessary) we should require that $t_z\ln z$ is about a constant, say $t_z = 1/\ln z$ for $z > e$. Substituting this into (7.33) yields

$$h'\left(1/\ln z\right) = -z\ln z, \quad \text{or} \quad \text{with } s = 1/\ln z, \quad h'(s) = -\frac{e^{1/s}}{s}.$$

Thus we may take

$$h(t) = \int_t^1 \frac{e^{1/s}}{s}ds,$$

and we obtain

$$f_M(t) = K\exp\left(-\int_{1-t}^1 \frac{e^{1/s}}{s}ds\right), \quad 0 < t < 1, \quad \text{where} \quad K^{-1} = \int_0^1 e^{-h(u)}du.$$

(Note that t_z is indeed the local minimum of (7.32).) A graph of the density f_M is given in Figure 5.

For the lower bound, as

$$p_\delta = K\int_{1-\delta}^1 e^{-h(1-t)}dt = K\int_0^\delta e^{-h(t)}dt,$$

we obtain

$$\ln\mathbb{P}(R \ge x) = \frac{\ln(1-c)}{\ln(1-cq/x)}\ln\left(K\int_0^{cq/x} e^{-h(t)}dt\right) \sim -\frac{\ln(1-c)}{cq}x\ln\left(\int_0^{cq/x} e^{-h(t)}dt\right).$$

We need the following lemma which we justify below.

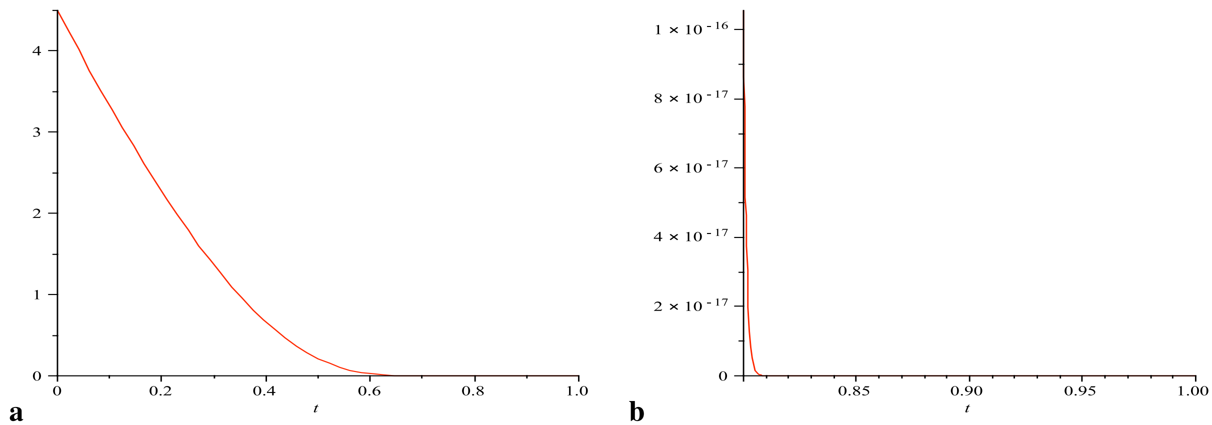

Figure 5: (a) The density f_M and (b) its detail closer to 1.

Lemma 8.

$$\frac{y \ln\left(\int_0^{1/y} e^{-h(t)} dt\right)}{e^y} \to -1, \quad \textit{as} \quad y \to \infty. \tag{7.34}$$

Using this lemma with $y = x/(cq)$ and $c = B/q$ we get, asymptotically,

$$\frac{\ln \mathbb{P}(R \geq x)}{e^{x/B}} \geq -\frac{\ln(1 - B/q)}{Be^{x/B}} x \ln\left(\int_0^{B/x} e^{-h(t)} dt\right) \sim \ln(1 - B/q),$$

which implies (7.29). □

Proof of Lemma 8: We re-write the left–hand side of (7.34) as

$$\frac{\ln\left(\int_0^{1/y} e^{-h(t)} dt\right)}{e^y/y},$$

and apply d'Hospital rule. The first differentiation gives

$$\frac{(-1/y^2)e^{-h(1/y)}}{\left(\int_0^{1/y} e^{-h(t)} dt\right)(-e^y/y^2 + e^y/y)} = \frac{e^{-h(1/y)} e^{-y}/(1-y)}{\int_0^{1/y} e^{-h(t)} dt}.$$

Differentiating again we get

$$\frac{(1/y^2)h'(1/y)e^{-h(1/y)}e^{-y}/(1-y) + e^{-h(1/y)} \frac{d}{dy}\left(e^{-y}/(1-y)\right)}{(-1/y^2)e^{-h(1/y)}} = -h'\left(\frac{1}{y}\right)\frac{e^{-y}}{1-y} - y^2 \frac{d}{dy}\left(\frac{e^{-y}}{1-y}\right).$$

Since $h'(s) = -e^{1/s}/s$ the first term goes to -1 as $y \to \infty$ while the second is $o(1)$. □

References

[1] G. Alsmeyer, A. Iksanov, and U. Rösler. On distributional properties of perpetuities. *J. Theoret. Probab.* to appear, 2008. http://www.citebase.org/abstract?id=oai:arXiv.org:0803.3716.

[2] M. Białkowski and J. Wesołowski. Asymptotic behavior of some random splitting schemes. *Probab. Math. Statist.*, 22:181–191, 2002.

[3] L. Bondesson. A general result on infinite divisibility. *Ann. Probab.*, 7(6):965–979, 1979.

[4] M. Brown. Error bounds for exponential approximations of geometric convolutions. *Ann. Probab.*, 18(3):1388–1402, 1990.

[5] J.-F. Chamayou and G. Letac. Explicit stationary distributions for compositions of random functions and products of random matrices. *J. Theoret. Probab.*, 4:3–36, 1991.

[6] P. Embrechts and C. M. Goldie. Perpetuities and random equations. In *Asymptotic statistics (Prague, 1993)*, Contrib. Statist., pages 75–86. Physica, Heidelberg, 1994.

[7] C. M. Goldie. Implicit renewal theory and tails of solutions of random equations. *Ann. Appl. Probab.*, 1(1):126–166, 1991.

[8] C. M. Goldie and R. Grübel. Perpetuities with thin tails. *Adv. in Appl. Probab.*, 28:463–480, 1996.

[9] D. R. Grey. Regular variation in the tail behaviour of solutions of random difference equations. *Ann. Appl. Probab.*, 4:169–183, 1994.

[10] A. K. Grincevičjus. On a limit distribution for a random walk on lines. *Litovsk. Mat. Sb.*, 15:79–91, 243, 1975.

[11] P. Hitczenko and G. S. Medvedev. Bursting oscillations induced by small noise, 2007. http://www.citebase.org/abstract?id=oai:arXiv.org:0712.4074.

[12] Z. J. Jurek. Selfdecomposability perpetuity laws and stopping times. *Probab. Math. Statist.*, 19:413–419, 1999.

[13] H. Kesten. Random difference equations and renewal theory for products of random matrices. *Acta Math.*, 131:207–248, 1973.

[14] M. Knape and R. Neininger. Approximating perpetuities. *Methodology and Computing in Applied Probability* to appear, 2008. http://www.citebase.org/abstract?id=oai:arXiv.org:0711.1099.

[15] M. A. Krasnosel′skiĭ and Ja. B. Rutickiĭ. *Convex functions and Orlicz spaces*. Translated from the first Russian edition by Leo F. Boron. P. Noordhoff Ltd., Groningen, 1961.

[16] G. Letac. A contraction principle for certain Markov chains and its applications. In *Random matrices and their applications (Brunswick, Maine, 1984)*, number 50 in Contemp. Math., pages 263–273. Amer. Math. Soc., Providence, RI., 1986.

[17] O. Thorin. On the infinite divisibility of the lognormal distribution. *Scand. Actuar. J.*, (3):121–148, 1977.

[18] O. Thorin. On the infinite divisibility of the Pareto distribution. *Scand. Actuar. J.*, (1):31–40, 1977.

[19] W. Vervaat. On a stochastic difference equation and a representation of nonnegative infinitely divisible random variables. *Adv. in Appl. Probab.*, 11(4):750–783, 1979.

[20] N. Yannaros. Randomly observed random walks. *Comm. Statist. Stochastic Models*, 7(2):219–231, 1991.